# UNIVERSITY OF OSLO DEPARTMENT OF ECONOMICS 

Exam: ECON4160 - Econometrics - Modeling and systems estimation
Date of exam: Friday, May 20, 2005 Grades will be given: Monday, June 13
Time for exam: 2:30 p.m. - 5:30 p.m.
The problem set covers 5 pages
Resources allowed:

- All printed and written resources, as well as calculator

Please answer both part I) and part II) of the problem set. Both parts will be given equal weight in the evaluation.

The grades given: A-F, with A as the best and E as the weakest passing grade. F is fail.

## Part I)

Enclosed are some results from estimations of the relative demand for different types of labour in the economy. The labour force is classified into three types of labour, according to the level of education. We have observations of relative employment and relative wages from 12 different European countries from the years 1974 to 2001. We do not have observations from each country each year, and the total number of observations of country $x$ year is 131 .

Model 1 of the appendix reports estimation results from a seemingly unrelated regression model with two equations, explaining the relative demand for two different types of workers: relempl1 = log of relative employment for group 1 relative to group 3, relempl2= log of relative employment for group 2 relative to group 3. The unit of observation is country per year. The equation for relempl1 includes a measure of relative wages for group 1 :
relwage1= the log of wages for group 1 relative to the wages of group 3, and similarly the equation for relemp2 includes a measure of relative wages for group 2: relwage $2=$ the log of wages for group 2 relative to the wage of group 3 .

Both relations include in addition the same set of 11 different country dummies (of which only ctryechp1,ctryechp2 and ctryechp3 are shown in the table) and 12 country specific time trends (of which only tnor, tspa and tswe are shown in the table), as well as a common time trend squared ( $t 2$ ). The table also reports the correlation matrix of residuals between the two equations.

## Question 1.

Assume that relative wages are exogenous in this model. Set up the stochastic model and describe the assumptions underlying this estimation strategy (the Seemingly Unrelated Regression model). What is the advantage of using SUR rather than separate ordinary least square models (OLS) under these assumptions? Explain. Discuss the usefulness of the Breusch-Pagan test of independence of the residuals, whose test statistic is reported at the bottom of the page, for the choice of modeling strategy.

## Question 2.

Model II provides estimation results for a different, but related SUR model. This time the relative wage of both groups are included as regressors in both equations. Consider now running separate OLS regressions on the two equations in this model. Based on the information on this page (Model II), indicate what the OLS estimate would be for each of the two relative wage measures, relwage1 and relwage2, in the OLS regression of relempl2.
Explain how you arrived at these numbers. In this model, all parameters are allowed to vary freely. What type of parameter restrictions would lead you to prefer SUR rather than OLS?

## Question 3.

Assume now that the correlation between the disturbance terms in the two equations is zero. Assume furthermore that the disturbance term, $\mathrm{u}_{2}$, of the second equation is homoskedastic, $\left[\operatorname{var}\left(\mathrm{u}_{2 \mathrm{i}}\right)=\sigma_{2}^{2}\right]$, but that the variance of the disturbance term, $\mathrm{u}_{1}$, of the first equation, is proportional to the inverse of the size of the population in the country $\left[\operatorname{var}\left(\mathrm{u}_{1 \mathrm{i}}\right)=\sigma_{1}{ }^{2} / \mathrm{POP}_{\mathrm{i}}\right]$. Suggest an estimation method in this case and compare it to using two separate OLS regressions.

## Part II)

## Question 4.

Consider the following model:

1) $y_{1 i}=a_{1}+b_{11} y_{2 i}+\gamma_{11} x_{1 i}+\gamma_{12} x_{2 i}+\gamma_{13} X_{3 i}+\gamma_{14} \mathrm{X}_{4 i}+v_{1 i}, \quad i=1, \ldots, n$
2) $y_{2 i}=a_{2}+b_{21} y_{1 i}+\gamma_{21} x_{1 i}+\gamma_{22} X_{2 i}+\gamma_{23} X_{3 i}+\gamma_{24} \mathrm{X}_{4 i}+v_{2 i}, \quad i=1, \ldots, n$

Discuss assumptions required for the x's to be exogenous in this model. Derive expressions for the expectation and variance, both conditional on the vector $x$, of the endogenous variables $y_{1}$ and $y_{2}$ in the model. Suggest exclusion restrictions that would make both equations exactly identifiable. Derive the indirect least square estimators (ILS) of $b_{11}$ and $b_{21}$ in that case.

## Question 5.

Consider the following simpler model:

1) $y_{1 i}=a_{1}+b_{11} y_{2 i}+\gamma_{11} x_{1 i}+\gamma_{12} x_{2 i}+\gamma_{13} x_{3 i}+\gamma_{14} x_{4 i}+v_{1 i}, \quad i=1, \ldots, n$
2) $y_{2 i}=a_{2}+b_{21} y_{1 i}+v_{2 i}, \quad i=1, \ldots, n$
where again the x 's are assumed to be exogenous, while $\mathrm{y}_{1}$ and $\mathrm{y}_{2}$ are endogenous.

Discuss identification of equation (2). Two researchers suggest the following instrumental variable estimators for $b_{21}$ each:
3) $\beta^{\mathrm{IV} 1}=M\left(y_{2}, x_{1}\right) / M\left(y_{1}, x_{1}\right)$
4) $\beta^{\mathrm{IV} 2}=\mathrm{M}\left(\mathrm{y}_{2}, \mathrm{x}_{2}\right) / \mathrm{M}\left(\mathrm{y}_{1}, \mathrm{x}_{2}\right)$
where $M(y, x)$ is the empirical covariance between $y$ and $x$. Show that both of these estimators are consistent. Still, these two estimators generally provide us with two different estimates. Suggest a statistics that can be used to evaluate if one of the two estimators, (3) or (4), is better than the other. Describe the optimal instrument for the estimation of $b_{21}$ in equation 2 . Under what restrictions would the IV-estimator given in equation (3) be the optimal one? Compare this estimator to the ILS estimator in that case.

## MODEL I)

Seemingly unrelated regression

| Equation | Obs | Parms | RMSE | "R-sq" | chi2 | P |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| relempl1 | 131 | 32 | . 0390764 | 0.9975 | 54832.97 | 0.0000 |
| relempl2 | 131 | 32 | . 0577976 | 0.9952 | 27647.58 | 0.0000 |


|  | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| relempl1 |  |  |  |  |  |  |
| ctryechp1 | -1.073731 | . 480049 | -2.24 | 0.025 | -2.01461 | -. 132852 |
| ctryechp2 | -. 8686135 | . 6285343 | -1.38 | 0.167 | -2.100518 | . 3632911 |
| ctryechp3 | -3.251651 | . 4527519 | -7.18 | 0.000 | -4.139028 | -2.364274 |
| tnor | . 2425992 | . 011017 | 22.02 | 0.000 | . 2210062 | . 2641923 |
| tspa | . 2614227 | . 0184553 | 14.17 | 0.000 | . 225251 | . 2975944 |
| tswe | . 2352761 | . 011219 | 20.97 | 0.000 | . 2132873 | . 2572649 |
| t2 | -. 0042338 | . 0002966 | -14.27 | 0.000 | -. 0048151 | -. 0036524 |
| relwage1 | -1.842812 | . 2979266 | -6.19 | 0.000 | -2.426737 | -1.258887 |
| _cons | -2.660008 | . 262904 | -10.12 | 0.000 | -3.17529 | -2.144725 |


| relempl2 \| |  |  |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| ctryechp1 \| | .056858 | .718363 | 0.08 | 0.937 | -1.351108 | 1.464824 |
| ctryechp2 \| | -2.946798 | .9518112 | -3.10 | 0.002 | -4.812314 | -1.081282 |
| ctryechp3 \| | -4.111755 | .7272126 | -5.65 | 0.000 | -5.537065 | -2.686444 |
| $\ldots \ldots . . \ldots .$. |  |  |  |  |  |  |
| tnor \| | .1385451 | .0160861 | 8.61 | 0.000 | .1070169 | .1700733 |
| tspa \| | .1508126 | .0276108 | 5.46 | 0.000 | .0966965 | .2049288 |
| tswe \| | .1237072 | .0164715 | 7.51 | 0.000 | .0914237 | .1559907 |
| t2 \| | -.0020179 | .0004366 | -4.62 | 0.000 | -.0028737 | -.0011621 |
| relwage2 \| | -5.113312 | 1.043669 | -4.90 | 0.000 | -7.158865 | -3.067759 |
| _cons \| | .1531386 | .410832 | 0.37 | 0.709 | -.6520773 | .9583545 |

Correlation matrix of residuals:
relempl1 relempl2
relempl1 1.0000
relempl2 $0.7421 \quad 1.0000$
Breusch-Pagan test of independence: chi2(1) $=$ 72.134, $\operatorname{Pr}=0.0000$

## MODEL II)

Seemingly unrelated regression

| Equation | Obs Parms |  | RMSE " | 'R-sq" | chi2 | P |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| relempl1 | 131 | 33.03 |  | 0.9976 | 54854.260 | 0.0000 |
| relempl2 | 131 | 33 . 05 |  | 0.9957 | 30164.950 | 0.0000 |
|  |  | Std. Err | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Conf. Interval] |  |
| relempl1 <br> ctryechp1 <br> ctryechp2 \| <br> ctryechp3 |  |  |  |  |  |  |
|  | -. 8623758 | . 5687173 | -1.52 | 0.129 | -1.977041 | . 2522897 |
|  | -1.256032 | . 6450218 | -1.95 | 0.052 | -2.520251 | . 0081874 |
|  | -3.857581 | . 5194508 | -7.43 | 0.000 | - 4.875686 | -2.839476 |
| tnor \| | . 2495834 | . 0114287 | 21.84 | 0.000 | . 2271836 | . 2719833 |
| tspa \| | . 2589843 | . 0185685 | 13.95 | 0.000 | . 2225907 | . 2953779 |
| tswe | . 2414091 | . 0116127 | 20.79 | 0.000 | . 2186487 | . 2641696 |
| t2 | -. 0043987 | . 0003027 | -14.53 | 0.000 | -. 0049919 | -. 0038055 |
| relwage1 | -2.631374 | . 588853 | -4.47 | 0.000 | -3.785505 | -1.477243 |
| relwage2 | -. 1363279 | 1.377138 | -0.10 | 0.921 | -2.835469 | 2.562813 |
| _cons | -2.114946 | . 31238 | -6.77 | 0.000 | - -2.7272 | -1.502693 |
| relempl2 \| |  |  |  |  |  |  |
| ctryechp1 | 1.327777 | . 8158387 | 1.63 | 0.104 | -. 2712376 | 2.926791 |
| ctryechp2 | -3.409934 | . 9252992 | -3.69 | 0.000 | -5.223487 | -1.59638 |
| ctryechp3 | -4.73703 | . 7451646 | -6.36 | 0.000 | - -6.197526 | -3.276534 |
| tnor | . 1596685 | . 0163948 | 9.74 | 0.000 | . 1275353 | . 1918016 |
| tspa | . 1403827 | . 026637 | 5.27 | 0.000 | . 0881752 | . 1925903 |
| tswe | . 1433452 | . 0166587 | 8.60 | 0.000 | . 1106948 | . 1759957 |
| t2 | -. 0024761 | . 0004342 | -5.70 | 0.000 | - -. 0033271 | -. 0016251 |
| relwage2 | -1.504857 | 1.975538 | -0.76 | 0.446 | -5.376839 | 2.367126 |
| relwage1 | -2.90679 | . 8447238 | -3.44 | 0.001 | -4.562418 | -1.251162 |
| _cons \| | . 9463941 | . 4481166 | 2.11 | 0.035 | . 0681017 | 1.824686 |

Correlation matrix of residuals:
relempl1 relempl2
relempl1 1.0000
relempl2 $0.7770 \quad 1.0000$
Breusch-Pagan test of independence: chi2(1) = 79.079, $\operatorname{Pr}=0.0000$

