UNIVERSITY OF OSLO DEPARTMENT OF ECONOMICS

Exam: ECON4160 – Econometrics – Modeling and systems estimation

Date of exam: Monday, May 26, 2008 Grades will be given: Thursday, June 12

Time for exam: 2:30 p.m. – 5:30 p.m.

The problem set covers 8 pages

Resources allowed:

• All written and printed resources, as well as calculators, are allowed

The grades given: A-F, with A as the best and E as the weakest passing grade. F is fail.

ECON 4160: ECONOMETRICS – MODELLING AND SYSTEMS ESTIMATION PROBLEM SET, EXAM SPRING 2008

PROBLEM 1 (weight: 60%)

We are interested in analyzing, from micro data, the relationship between female labour supply, measured as the actual number of hours worked per year, and the length of education and work experience, measured in years. To explore this a cross-section data set from 753 females in the US observed in 1975 for the following six variables has been compiled:

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Y1 = Number of hours worked in the year 1975
Y2 = Education, in years
X1 = Work experience, in years
Z1 = Father's education, in years
Z2 = Mother's education, in years
Z3 = Husband's education, in years
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We assume that (Y1, Y2) are endogenous variables, that X1 is exogenous, and that (Z1, Z2, Z3) have been proposed as candidates for being instruments for Y2 in the equation

(*)
$$Y1 = \alpha + \beta Y2 + \gamma X1 + U,$$

where U is a disturbance.

The estimation results and other printouts referred to below are obtained from PcGive and are given at the end of the problem.

(A): Give a stochastic specification of the model, and give reasons why treating (Y1,Y2) as jointly endogenous variables may be reasonable. In EQ(1)–EQ(2) two versions of (*) are estimated, the first with γ set to zero *a priori*, the second with both coefficients free. Explain briefly why the two equations give different estimates of β and why both sets of OLS estimates are inconsistent.

(B): Assume that (X1,Z1,Z2,Z3) are exogenous variables in the model to which (*) belongs and that the number of other equations and of other endogenous variables, say $N_*(\geq 1)$, is unknown. Show, by using the order condition, that (*) is identified regardless of the value of N_* .

(C): Consider the estimates in printouts EQ(3) and EQ(6). Explain briefly the terms 'IVE' and 'Additional instruments' and explain why the estimates are both consistent under the given assumptions. Why do they differ when computed from the 753 observations?

(D): Could X1 alone have served as an instrument for Y2 in (*)? State briefly the reason for your answer.

(E): Let (Q1, Q2, Q3, Q4) be four derived variables defined and calculated by PcGive by

Algebra code for variable transformations:

Q1 = X1+Z1; Q2 = X1+2*Z1; Q3 = X1+Z2; Q4 = X1+2*Z2;

Explain why (Q1,Q2,Q3,Q4) are all valid instruments for (*) and why the estimates in EQ(4)-EQ(5) coincide with those in EQ(3) and why the estimates in EQ(7)-EQ(8) coincide with those in EQ(6).

[Hint: Note that (i) both (X1,Q1) and (X1,Q2) are one-to-one (non-singular) transformations of (X1,Z1) and (ii) (X1,Q3) and (X1,Q4) are one-to-one (non-singular) transformations of (X1,Z2).]

(F): Explain briefly the estimation method used for equations EQ(9)-EQ(10), in particular how it differs from the methods used for equations EQ(3) and EQ(6). Which conclusions do you draw from printouts EQ(11)-EQ(12) and the correlation matrix below about the quality of the instruments? What would you conclude about the effect on the female labour supply of (a) a one year increase in education, (b) a one year increase in work experience?

(G): Good arguments may be given for treating X1 as an endogenous variable, jointly determined with Y1 and Y2. If you accept this, how would you then modify your model and proceed to estimate the coefficients of (*)? Explain briefly.

Means Y 740.5	71 58 12.	Y2 .287 10	X1 0.631	Z1 8.8088	Z2 9.2510	Z3 12.491
Standard Y 871.3	deviations 71 31 2.2	(using T-1) Y2 2802 8.	X1 0691	Z1 3.5723	Z2 3.3675	Z3 3.0208
Correlati Y1 1 Y2 0. X1 0. Z1 0.0 Z2 0.0 Z3 -0.00	ion matrix: Y1 1.0000 .10596 .40496 C 013671 057864 096504	Y2 0.10596 1.0000 0.066256 0.44246 - 0.43534 - 0.61195 -	X1 0.40496 0.066256 1.0000 0.078802 0.082179 0.036301	Z1 0.013671 0.44246 -0.078802 1.0000 0.57307 0.36670	Z2 0.057864 0.43534 -0.082179 0.57307 1.0000 0.32447	Z3 -0.0096504 0.61195 -0.036301 0.36670 0.32447 1.0000

PCGIVE PRINTOUTS FOR PROBLEM 1

MEANS, STANDARD DEVIATION AND CORRELATIONS. THE SAMPLE IS: 1 TO 753

EQ(1) Modelling Y1 by OLS-CS. The estimation sample is: 1 to 753 Coefficient Std.Error t-value t-prob Part.R^2 Y2 40.4890 13.87 2.92 0.004 0.0112 Constant 243.094 173.3 1.40 0.161 0.0026 sigma R^2 866.986 RSS 564499772 0.0112276 F(1,751) =8.528 [0.004]** log-likelihood -6161.52 DW 0.973 753 no. of parameters no. of observations 2 740.576 var(Y1) mean(Y1) 758180 ********** EQ(2) Modelling Y1 by OLS-CS. The estimation sample is: 1 to 753 Coefficient Std.Error t-value t-prob Part.R^2 2.38 0.0075 Y2 12.74 30.3699 0.017 12.0 43.1593 3.599 0.000 0.1609 X1 -91.3922 Constant 161.3 -0.567 0.571 0.0004 794.728 RSS 473694833 sigma F(2,750) =0.170281 76.96 [0.000]** R^2 log-likelihood 1.18 DŴ -6095.49 753 no. of parameters 740.576 var(Y1) no. of observations 3 758180 mean(Y1) EQ(3) Modelling Y1 by IVE-CS. The estimation sample is: 1 to 753 Coefficient Std.Error t-value t-prob 1.37 v Y2 38.9067 28.31 0.170 X1 42.9995 3.632 11.8 0.000 Constant -194.584345.5 -0.563 0.574 794.966 RSS 473978536 sigma Reduced form sigma 796.74 753 no. of parameters s 2 no. of instruments no. of observations 3 no. endogenous variables 3 740.576 var(Y1) 758180 mean(Y1) Additional instruments: [0] = Z1 EQ(4) Modelling Y1 by IVE-CS. The estimation sample is: 1 to 753 Coefficient Std.Error t-value t-prob 38.9067 42.9995 1.37 11.8 Y2 Y 28.31 0.170 X1 3.632 0.000 345.5 -0.563 Constant -194.5840.574 794.966 RSS 473978536 sigma Reduced form sigma 796.74 753 no. of parameters les 2 no. of instruments 740.576 var(Y1) no. of observations 3 no. endogenous variables 3 758180 mean(Y1) Additional instruments: [0] = Q1

EQ(5) Modelling Y1 by IVE-CS. The estimation sample is: 1 to 753 Coefficient Std.Error t-value t-prob Y2 Y 38.9067 28.31 1.37 0.170 X1 42.9995 3.632 11.8 0.000 Constant -194.584345.5 -0.563 0.574 794.966 RSS 473978536 sigma Reduced form sigma 796.74 no. of observations 753 no. of parameters s 2 no. of instruments 3 no. endogenous variables З 740.576 var(Y1) 758180 mean(Y1) Additional instruments: [0] = Q2********** EQ(6) Modelling Y1 by IVE-CS. The estimation sample is: 1 to 753 Coefficient Std.Error t-value t-prob 79.0119 42.2486 Y2 0.007 Y 29.01 2.72 X1 3.667 11.5 0.000 -1.92 Constant -679.367354.0 0.055 802.418 RSS 482905547 sigma Reduced form sigma 793.73 753 no. of parameters les 2 no. of instruments 740.576 var(Y1) 3 no. of observations no. endogenous variables З 758180 mean(Y1) Additional instruments: [0] = Z2EQ(7) Modelling Y1 by IVE-CS. The estimation sample is: 1 to 753 Coefficient Std.Error t-value t-prob Y2 Y 79.0119 29.01 2.72 0.007 X1 42.2486 3.667 11.5 0.000 Constant -679.367 354.0 -1.92 0.055 802.418 RSS 482905547 sigma Reduced form sigma 793.73 753 no. of parameters s 2 no. of instruments no. of observations 3 no. endogenous variables 3 740.576 var(Y1) 758180 mean(Y1) Additional instruments: [0] = Q3******* EQ(8) Modelling Y1 by IVE-CS. The estimation sample is: 1 to 753 Coefficient Std.Error t-value t-prob 79.0119 29.01 2.72 0.007 79.0119 42.2486 ¥2 Y 29.01 2.72 3.667 0.000 X1 11.5 -1.92 Constant -679.367 354.0 0.055 802.418 RSS 482905547 sigma Reduced form sigma 793.73 753 no. of parameters 2 no. of instruments no. of observations 3 no. endogenous variables З 740.576 var(Y1) 758180 mean(Y1) Additional instruments: [0] = Q4******

EQ(9) Modelling Y1 by IVE-CS. The estimation sample is: 1 to 753 Coefficient Std.Error t-value t-prob Y2 Y 22.6733 18.74 1.21 0.227 X1 43.3034 3.610 12.0 0.000 Constant 1.64213 231.5 0.00709 0.994 794.922 RSS 473925434 sigma Reduced form sigma 794.47 753 no. of parameters s 2 no. of instruments no. of observations 3 no. endogenous variables 5 740.576 var(Y1) 758180 mean(Y1) Additional instruments: [0] = Z1[1] = Z2[2] = Z3EQ(10) Modelling Y1 by IVE-CS. The estimation sample is: 1 to 753 Coefficient Std.Error t-value t-prob Y2 Y 22.6733 43.3034 18.74 $1.21 \\ 12.0$ 0.227 3.610 X1 0.000 231.5 0.00709 1.64213 0.994 Constant 794.922 RSS sigma 473925434 Reduced form sigma 794.47 753 no. of parameters 2 no. of instruments no. of observations З no. endogenous variables 5 740.576 var(Y1) 758180 mean(Y1) Additional instruments: [0] = Q1 [1] = Q3 [2] = Z3 ****** EQ(11) Modelling Y1 by OLS-CS. The estimation sample is: 1 to 753 Coefficient Std.Error t-value t-prob Part.R² 0.1692 X1 44.5060 3.605 12.3 0.000 Ζ1 -0.56844210.18 -0.0558 0.956 0.0000 Z2 26.3366 10.63 2.48 0.013 0.0081 Z3 -7.74773 10.43 -0.743 0.458 0.0007 Constant 125.588 138.8 0.905 0.366 0.0011 sigma R^2 794.472 472127057 RSS 0.173027 F(4,748) =39.13 [0.000]** log-likelihood -6094.24 DW 1.17 no. of parameters no. of observations 753 5 740.576 var(Y1) 758180 mean(Y1) EQ(12) Modelling Y2 by OLS-CS. The estimation sample is: 1 to 753 t-prob Part.R² Coefficient Std.Error t-value 0.007590 0.000 0.0230 X1 0.0318243 4.19 0.101756 0.02144 4.75 Z1 0.000 0.0292 Z2 0.02238 5.83 0.0434 0.130410 0.000 Z3 0.373721 0.02195 17.0 0.000 0.2793 Constant 5.17748 0.2921 17.7 0.000 0.2957 1.6726 sigma RSS 2092.60733 0.464812 F(4,748) = -1453.28 DW R^Ž 162.4 [0.000]** log-likelihood 2 no. of observations 753 no. of parameters 5

mean(Y2)

PROBLEM 2 (weight: 40%)

(A): Consider the simple macro model

(1)
$$C_t = \alpha + \beta Y_t + u_t$$

$$Y_t = C_t + I_t + G_t.$$

where Y_t (= GNP) and C_t (= Total Private Consumption) are endogenous, I_t (= Total Gross Investment) and G_t (= Total Public Expenditure) are exogenous variables, and u_t is a disturbance. Complete the model description and explain which of its equations can be identified from time series on (Y_t, C_t, I_t, G_t) . The marginal propensity to consume, β , can be estimated consistently by instrumental variables in four different ways, by using as instruments for Y_t , respectively, (i) only I_t , (ii) only G_t , (iii) $I_t + G_t$, or (iv) both I_t and G_t . Which of alternatives (i)–(iv) would you prefer if you believe in this simple model? State the reason for your answer.

(B): An extended version of the macro model is also of interest:

 $C_t = \alpha_1 + \beta_1 Y_t + u_t,$

(4)
$$I_t = \alpha_2 + \beta_2 (Y_t - Y_{t-1}) + \gamma_2 G_t + v_t$$

where (4), with $\beta_2 > 0$, $\gamma_2 > 0$, represents a hypothesis that gross investment responds partly to the increase in GNP and partly to certain components of Total Public Expenditure, and v_t is a disturbance. Complete the model description also in this case. Decide which of the model's equations can be identified from time series of (Y_t, C_t, I_t, G_t) . How would you now estimate the consumption function?

[Hint: In interpreting (4) and specifying the model stochastically, you may consider it as having the form

$$I_t = \alpha_2 + \beta_2 Y_t + \beta_3 Z_t + \gamma_2 G_t + v_t$$

with the linear restriction $\beta_3 = -\beta_2$ imposed and with $Z_t = Y_{t-1}$ considered as predetermined (with properties which in this context can be treated as coinciding with those of an exogenous variable).]

(C): The reduced form equation for Y_t , obtained by inserting (3) and (4) into the national budget identity (5) and solving for Y_t is (derivation not required)

(6)
$$Y_t = a + bG_t + cY_{t-1} + \varepsilon_t,$$

where

$$a = \frac{\alpha_1 + \alpha_2}{1 - \beta_1 - \beta_2}, \quad b = \frac{1 + \gamma_2}{1 - \beta_1 - \beta_2}, \quad c = -\frac{\beta_2}{1 - \beta_1 - \beta_2}, \quad \varepsilon_t = \frac{u_t + v_t}{1 - \beta_1 - \beta_2}.$$

Would you consider (6) as describing a lag distribution, and if so which form does it have? Assume that consistent estimates of $(\beta_1, \beta_2, \gamma_2)$, satisfying $\beta_2 < \frac{1}{2}(1-\beta_1) \Longrightarrow |c| < 1$, are available (you are not required to propose an estimation procedure). Explain how you from this information would estimate b and c consistently and explain briefly how you from the estimates obtained, symbolized by $\hat{}$, would proceed to compute the effect of a one unit increase in G in a particular year

- (a) on Y in the current year,
- (b) on Y in the next year, and
- (c) on Y in the long run, i.e., the sum of the effects in the current and all future years.

[Hint: To illustrate your points you may well use numerical values, say

 $(\hat{\beta}_1, \hat{\beta}_2, \hat{\gamma}_2) = (0.65, 0.1, 0.05) \implies (\hat{b}, \hat{c}) = (4.2, -0.4).$]