

# UNIVERSITY OF OSLO

## Faculty of Mathematics and Natural Sciences

Examination in: STK4011/STK9011 — Statistical inference theory

Day of examination: Tuesday 8 December 2015.

Examination hours: 09.00–13.00.

This problem set consists of 3 pages.

Appendices: None.

Permitted aids: “List of formulas for STK1100 and STK1110”  
and “Summary of results for STK4011/9011”.

Please make sure that your copy of the problem set is complete before you attempt to answer anything.

### Problem 1

Let  $X_1, \dots, X_n$  be independent Bernoulli variables with success probability  $p$ . So for each  $i = 1, \dots, n$ , we have  $P(X_i = 1) = 1 - P(X_i = 0) = p$ .

- Show that  $T = \sum_{i=1}^n X_i$  is a sufficient statistic and that the family of pmfs of  $T$  has a monotone likelihood ratio.
- Find the uniform most powerful test of  $H_0 : p \leq p_0$  versus  $H_1 : p > p_0$  and describe how you may determine the size of the test.
- Derive the likelihood ratio test of  $H_0 : p = p_0$  versus  $H_1 : p \neq p_0$  and describe how you may use the test to find an approximate  $1 - \alpha$  confidence interval for  $p$ .

### Problem 2

We assume that  $X \sim \text{gamma}(\alpha, \beta)$ . So  $X$  has pdf

$$f(x|\alpha, \beta) = \begin{cases} \frac{1}{\beta^\alpha \Gamma(\alpha)} x^{\alpha-1} e^{-x/\beta} & \text{if } x > 0, \\ 0 & \text{if } x \leq 0. \end{cases}$$

- Show that

$$E(X^\nu) = \beta^\nu \frac{\Gamma(\alpha + \nu)}{\Gamma(\alpha)}.$$

For which values of  $\nu$  is the formula valid?

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We further assume that  $Y \sim \text{gamma}(\gamma, \beta)$ , and that  $X$  and  $Y$  are independent. Let  $U = X + Y$  and  $V = X/(X + Y)$ .

- b) Derive the joint pdf of  $(U, V)$ .
- c) Show that  $V \sim \text{beta}(\alpha, \gamma)$ .

### Problem 3

Let  $X_1, X_2, \dots, X_n$  be independent and identically distributed with pdf

$$f(x|\theta) = \begin{cases} \frac{\theta^3}{2} x^2 e^{-\theta x} & \text{if } x > 0, \\ 0 & \text{if } x \leq 0. \end{cases}$$

- a) Find a complete sufficient statistic for  $\theta$ .
- b) Find the maximum likelihood estimator  $\hat{\theta}$  of  $\theta$ .
- c) Show that

$$E_{\theta} \hat{\theta} = \frac{3n}{3n-1} \theta$$

and

$$\text{Var}_{\theta} \hat{\theta} = \frac{9n^2}{(3n-1)^2(3n-2)} \theta^2.$$

- d) Find the best unbiased estimator of  $\theta$  and show that its variance is larger than the Cramér-Rao lower bound.
- e) Is there a function of  $\theta$ , say  $\tau(\theta)$ , for which there exists an unbiased estimator whose variance attains the Cramér-Rao lower bound? If so, find it. If not, show why not.

### Problem 4

Let  $X_1, \dots, X_n$  be independent and identically distributed with pdf

$$f_X(x|\theta) = \begin{cases} \frac{\theta^3}{2} x^2 e^{-\theta x} & \text{if } x > 0, \\ 0 & \text{if } x \leq 0. \end{cases}$$

and let  $Y_1, \dots, Y_n$  be independent and identically distributed with pdf

$$f_Y(y|\phi) = \begin{cases} \frac{\phi^3}{2} y^2 e^{-\phi y} & \text{if } y > 0, \\ 0 & \text{if } y \leq 0. \end{cases}$$

Moreover, we assume that the  $X_i$ 's and  $Y_i$ 's are independent.

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- a) Find the likelihood ratio test statistic for testing of  $H_0 : \theta = \phi$  versus  $H_1 : \theta \neq \phi$ .

We introduce

$$T = \frac{\sum_{i=1}^n X_i}{\sum_{i=1}^n X_i + \sum_{j=1}^n Y_j}$$

- b) Show that the likelihood ratio test rejects  $H_0$  if  $T(1 - T) \leq k$  for a constant  $k$ .
- c) Explain that an alternative formulation of the likelihood ratio test in question b is to reject  $H_0$  if  $T \leq a$  or  $T \geq 1 - a$  for a constant  $a$ . Find the distribution of  $T$  when  $H_0$  is true and explain how you may use this distribution to find the value of  $a$  that gives a test with size  $\alpha$ .

**END**