

Lecture notes XII

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Experiments and Quasi-Experiments

SBW Ch. 13

Randomized controlled experiments randomly selects subjects from a population and then randomly assigns them either to a treatment group or control group

E.g. a drug is given to randomly selected individuals, while others get a "placebo" (harmless substitute)

Controlled experiments rare in economics. But similar methods used in program evaluation.

E.g. job training programs, R&D tax refund schemes, etc.

Let X_i be the treatment level,
e.g. a dummy variable: $Y_i = 1$ if treated, 0 else.

The effect of the treatment on the
outcome variable (Y_i) is given by a
regression equation:

① $Y_i = \beta_0 + \beta_1 X_i + u_i$

i) If treatment level, X_i , is assigned
completely at random \Rightarrow independent of
all other determinants of $Y_i \Rightarrow E(u_i | X_i) = 0$
 \Rightarrow no omitted variable bias

ii) The causal effect of the treatment
is $\Delta = E(Y_i | X_i = x) - E(Y_i | X_i = 0)$

$\Rightarrow \Delta = \beta_1 x$. Δ is called the
treatment effect.

Even in controlled experiments real-world experiments deviate from the idealized experiment, which may cause

$$E(u_i | X_i) \neq 0$$

Reasons for failure of complete randomization:

- i) The assignment and participation depend on the subject's preferences or motivation
- ii) Partial compliance \rightarrow a subject may be assigned to treatment but does not follow the instructions (e.g. does not take the drug)
- iii) Subjects may drop out after initially participating \rightarrow Attrition.
Attrition may be endogenous \Rightarrow bias estimates of β_1 using OLS

Regression based estimators of treatment effects

Let W_{i1}, \dots, W_{ik} denote ~~r~~ variables that characterize the subject i prior to the treatment, i.e. not affected by the treatment \rightarrow

(2)
$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 W_{i1} + \dots + \beta_{1+k} W_{ik} + u_i$$

- X_i is the treatment variable
- W_{ik} are the control variables

Control variables are not variables of interest per se, but are included to avoid omitted variable bias in the estimate of β_1 . May also lead to more efficient estimator of β_1 .

Standard OLS assumption: $E(u_i | X_i, W_{i1}, \dots, W_{ik}) = 0$,
i.e. conditional mean-zero assumption

Weaker assumption: $E(u_i | X_i, W_{i1}, \dots, W_{ik}) = E(u_i | W_{i1}, \dots, W_{ik})$, i.e. conditional mean independence

Example: Conditional random assignment.

That is, X_i may depend on the control variables, W_{ki} . However, given W_{i1}, \dots, W_{iK} , U_i is independent of X_i .

- E.g. job training programs distributed to persons with low education.

Differences-in-Differences Estimator

Experimental data are often panel data:

The same unit i is observed before and after the experiment.

Let ΔY_i be the change in the value of Y_i ~~after~~ after the treatment compared to before the treatment

③
$$\Delta Y_i = \beta_0 + \beta_1 X_i + U_i$$

The β_1 - estimate using OLS is the Differences-in-differences estimator (DID)

Can also include control variables W_{i1}, \dots, W_{iK} in ③

Main motivation for DID estimator is that the treatment may be correlated with pre-treatment value of Y_i .

The DID estimator removes fixed effects, d_i , that may be correlated with X_i . $\therefore \text{cov}(d_i, X_i) \neq 0$

$$Y_{1i} = \beta_{10} + d_i + u_{1i}$$

$$Y_{2i} = \beta_{20} + \beta_1 X_i + d_i + u_{2i}$$

$$\Delta Y_i = \beta_{20} - \beta_{10} + \beta_1 X_i + u_{2i} - u_{1i}$$

$$\equiv \beta_0 + \beta_1 X_i + u_i$$

Quasi-experiments

- randomness is introduced by variations in individual circumstances that make it appear as if the treatment is randomly assigned.

If the "as if" random variation only partly determines X_i , so that $\text{cov}(X_i, u_i) \neq 0$, the exogenous part of the variation in X_i may be identified using instrumental variables

Quasi-experimental estimates of the effect of class
size on achievement in Norway

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Introduction

- ▶ The effect of class size on achievement is still an unresolved issue in education research.
- ▶ Endogeneity problems may severely bias naive OLS-estimates of the class size effect
- ▶ Exogenous sources of variation in class size are key for a credible identification of the class size effect.
- ▶ Various recent studies acknowledge this and apply ingenious identification methods.
- ▶ Still no definite conclusion about the magnitude and even the sign of the class-size effect:
 - ▶ Recent experimental and quasi-experimental studies find negative effects (Angrist & Lavy 1999, Krueger 1999).
 - ▶ Some, which also exploit credible identification strategies, fail to confirm these findings (f.e. Hoxby 2000, Hanushek et al. 1996).

Table 2: Regression of actual class size on observables

	(1)	(2)
Girl	0.019 (0.029)	0.019 (0.026)
Age	-0.064 (0.051)	0.071 (0.046)
ln(family income)	0.430 (0.042)***	0.168 (0.025)***
Education mother (years)	0.047 (0.008)***	0.012 (0.007)*
Education father (years)	0.104 (0.009)***	0.036 (0.006)***
1st or 2nd generation immigrants	1.313 (0.155)***	-0.223 (0.138)
Parents non-cohabiting	0.244 (0.045)***	-0.017 (0.040)
ln(pop. size school district)		1.211 (0.068)***
ln(rural pop. size school district)		0.434 (0.148)***
Year=2002		0.097 (0.130)
adj. R-squared	0.0192	0.2103
N	115,084	115,084
N schools	1,000	1,000

Estimating class size effects

Assume that achievement of student i (y_i) is generated by the following equation:

$$y_i = x_i'\beta + w_{s(i)}'\alpha + \delta \cdot \overline{CS}_{s(i)} + \eta_{s(i)} + \psi_{t(i)} + \epsilon_i \quad (1)$$

- ▶ x_i is a vector of observable attributes of the student and his parents
- ▶ $w_{s(i)}$ a vector of observable school and teacher characteristics
- ▶ $\overline{CS}_{s(i)}$ is student i 's class size
- ▶ $\eta_{s(i)}$ a school effect
- ▶ $s(i)$ identifies the school of pupil i
- ▶ $\psi_{t(i)}$ is an effect for the year in which student i is in her final year of lower secondary school (2001/2 or 2002/3)
- ▶ ϵ_i is all other determinants of achievement such as unobserved attributes of the student, parents and community.
- ▶ The coefficient of interest is δ , the class size effect.

Test score

- ▶ Exam results from national tests measured in the end of lower secondary school, grade 9.
 - ▶ 2001/2002 and 2002/2003.
 - ▶ Mathematics, English, Norwegian1 (Hovedmål) and Norwegian2 (Sidemål)
- ▶ Each student is only examined in one of the subjects, decided centrally shortly before the exam
 - ▶ the students and their teachers cannot prepare the whole school year for a particular exam
- ▶ The grading scale goes from 1 to 6 where 1 is fail and 6 is top score.
- ▶ Average scores for each of the four subjects are around 3.5 with standard deviations almost equal to 1.

Table 3: OLS. Dependent variable is exam results in 9th grade, unit of observations is the individual student's exam result

	Mathematics			Languages (pooled)		
	(1)	(2)	(3)	(4)	(5)	(6)
Average class size grade 7-9	0.005 (0.003)*	-0.006 (0.002)**	-0.005 (0.002)**	0.004 (0.002)**	-0.006 (0.002)***	-0.005 (0.002)***
<i>Individual characteristics</i>						
Girl		0.018 (0.012)	0.018 (0.012)		0.506 (0.010)***	-0.506 (0.0010)***
Age (years)		0.071 (0.019)***	0.072 (0.019)***		0.037 (0.014)**	0.036 (0.014)**
ln(family income)		0.180 (0.009)***	0.181 (0.009)***		0.144 (0.008)***	0.144 (0.008)***
Education mother (years)		0.075 (0.002)***	0.075 (0.002)***		0.057 (0.002)***	0.057 (0.002)***
Education father (years)		0.066 (0.002)***	0.066 (0.002)***		0.049 (0.002)***	0.049 (0.002)***
1st or 2nd generation immigrants		-0.446 (0.032)***	-0.437 (0.032)***		-0.232 (0.025)***	-0.225 (0.025)***
Parents have different address		-0.345 (0.013)***	-0.342 (0.013)***		-0.189 (0.009)***	-0.188 (0.009)***
ln(pop. size school district)		0.017 (0.008)**	0.028 (0.008)***		0.011 (0.005)**	0.015 (0.006)**
ln(rural pop. size school district)		-0.018 (0.016)	-0.024 (0.016)		-0.011 (0.010)	-0.016 (0.010)
<i>School characteristics:</i>						
Average teacher education (years)			0.042 (0.055)			0.070 (0.042)*
Average teacher experience			0.016 (0.003)***			0.011 (0.002)***
Fract. of female teachers			-0.097 (0.110)			0.059 (0.077)
Fract. teachers with a temp. contract			0.067 (0.088)			0.022 (0.065)
Combined school			0.052 (0.027)**			0.028 (0.022)
R-squared	0.001	0.164	0.167	0.011	0.171	0.173
N pupils	38,045	38,045	38,045	77,039	77,039	77,039
N schools	755	755	755	958	958	958

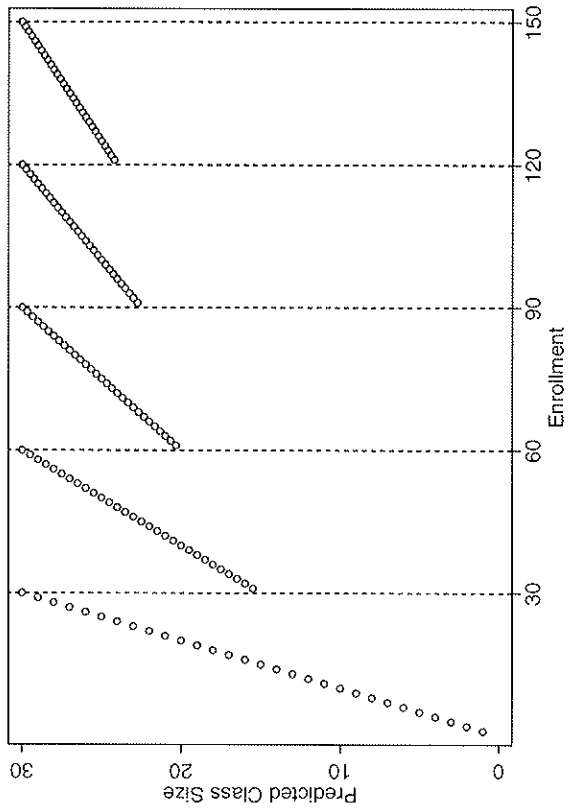
Note: Standard errors are heteroscedasticity robust and corrected for school level clustering. All regressions include a control for year of observation; the language regressions also include dummies for the language the student took the exam in.

1st approach: maximum class size rule as predictor for endogenous actual class size (AL, 1999)

- ▶ Lower secondary schools in Norway are subject to maximum class size rules of 30 students per class.
- ▶ This rule creates a discontinuous relation between enrollment and class size.
 - ▶ Just above multiples of 30 class size drops substantially.

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Rule



1st approach (cont'd)

- ▶ Predicted class size in grade 7 is used to instrument average class size in lower secondary school.
 - ▶ Enrollment (and therefore predicted class size) in subsequent grades is potentially endogenous

Results: teacher hours

	All	DS±5
		<i>First Stage</i> [†]
Predicted Class size	-0.671 (0.080)***	-0.413 (0.106)***
		<i>Second Stage - Mathematics</i>
Teacher Hours/Pupil	0.005 (0.006)	-0.012 (0.020)
		<i>Second Stage - Language</i>
Teacher Hours/Pupil	-0.002 (0.003)	-0.010 (0.007)

IV Regression with heterogeneous causal effects

Assume:

$$\begin{aligned}
 Y_i &= \beta_{0i} + \beta_{1i} X_i + u_i \\
 &= \beta_0 + \beta_1 X_i + (\beta_{0i} - \beta_0) + (\beta_{1i} - \beta_1) X_i + u_i \\
 &= \beta_0 + \beta_1 X_i + w_i, \quad w_i = \beta_{0i} - \beta_0 + (\beta_{1i} - \beta_1) X_i + u_i
 \end{aligned}$$

$$- E(w_i) = 0$$

$$- E(w_i | X_i) \neq 0$$

Assume Z_i is a valid instrument for X_i :

$$X_i = \pi_{0i} + \pi_{1i} Z_i + V_i$$

The TSLS estimator of β_1 is then

$$\hat{\beta}_1^{TSLS} = \frac{S_{ZY}}{S_{ZX}} \xrightarrow{P} \frac{\text{cov}(Z, Y)}{\text{cov}(Z, X)} = \frac{E(\beta_{1i} \pi_{1i})}{E(\pi_{1i})}$$

is called local average treatment effect (LATE),

while $E(\beta_{1i})$ is called average treatment effect.